

# Part IA — Probability

## Theorems

Based on lectures by R. Weber

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These notes are not endorsed by the lecturers, and I have modified them (often significantly) after lectures. They are nowhere near accurate representations of what was actually lectured, and in particular, all errors are almost surely mine.

### Basic concepts

Classical probability, equally likely outcomes. Combinatorial analysis, permutations and combinations. Stirling's formula (asymptotics for  $\log n!$  proved). [3]

### Axiomatic approach

Axioms (countable case). Probability spaces. Inclusion-exclusion formula. Continuity and subadditivity of probability measures. Independence. Binomial, Poisson and geometric distributions. Relation between Poisson and binomial distributions. Conditional probability, Bayes's formula. Examples, including Simpson's paradox. [5]

### Discrete random variables

Expectation. Functions of a random variable, indicator function, variance, standard deviation. Covariance, independence of random variables. Generating functions: sums of independent random variables, random sum formula, moments.

Conditional expectation. Random walks: gambler's ruin, recurrence relations. Difference equations and their solution. Mean time to absorption. Branching processes: generating functions and extinction probability. Combinatorial applications of generating functions. [7]

### Continuous random variables

Distributions and density functions. Expectations; expectation of a function of a random variable. Uniform, normal and exponential random variables. Memoryless property of exponential distribution. Joint distributions: transformation of random variables (including Jacobians), examples. Simulation: generating continuous random variables, independent normal random variables. Geometrical probability: Bertrand's paradox, Buffon's needle. Correlation coefficient, bivariate normal random variables. [6]

### Inequalities and limits

Markov's inequality, Chebyshev's inequality. Weak law of large numbers. Convexity: Jensen's inequality for general random variables, AM/GM inequality.

Moment generating functions and statement (no proof) of continuity theorem. Statement of central limit theorem and sketch of proof. Examples, including sampling. [3]

# Contents

<b>0</b>	<b>Introduction</b>	<b>3</b>
<b>1</b>	<b>Classical probability</b>	<b>4</b>
1.1	Classical probability . . . . .	4
1.2	Counting . . . . .	4
1.3	Stirling's formula . . . . .	4
<b>2</b>	<b>Axioms of probability</b>	<b>5</b>
2.1	Axioms and definitions . . . . .	5
2.2	Inequalities and formulae . . . . .	5
2.3	Independence . . . . .	5
2.4	Important discrete distributions . . . . .	6
2.5	Conditional probability . . . . .	6
<b>3</b>	<b>Discrete random variables</b>	<b>7</b>
3.1	Discrete random variables . . . . .	7
3.2	Inequalities . . . . .	8
3.3	Weak law of large numbers . . . . .	9
3.4	Multiple random variables . . . . .	9
3.5	Probability generating functions . . . . .	10
<b>4</b>	<b>Interesting problems</b>	<b>11</b>
4.1	Branching processes . . . . .	11
4.2	Random walk and gambler's ruin . . . . .	11
<b>5</b>	<b>Continuous random variables</b>	<b>12</b>
5.1	Continuous random variables . . . . .	12
5.2	Stochastic ordering and inspection paradox . . . . .	12
5.3	Jointly distributed random variables . . . . .	12
5.4	Geometric probability . . . . .	12
5.5	The normal distribution . . . . .	12
5.6	Transformation of random variables . . . . .	12
5.7	Moment generating functions . . . . .	13
<b>6</b>	<b>More distributions</b>	<b>14</b>
6.1	Cauchy distribution . . . . .	14
6.2	Gamma distribution . . . . .	14
6.3	Beta distribution* . . . . .	14
6.4	More on the normal distribution . . . . .	14
6.5	Multivariate normal . . . . .	14
<b>7</b>	<b>Central limit theorem</b>	<b>15</b>
<b>8</b>	<b>Summary of distributions</b>	<b>16</b>
8.1	Discrete distributions . . . . .	16
8.2	Continuous distributions . . . . .	16

## 0 Introduction

# 1 Classical probability

## 1.1 Classical probability

## 1.2 Counting

**Theorem** (Fundamental rule of counting). Suppose we have to make  $r$  multiple choices in sequence. There are  $m_1$  possibilities for the first choice,  $m_2$  possibilities for the second etc. Then the total number of choices is  $m_1 \times m_2 \times \cdots m_r$ .

## 1.3 Stirling's formula

**Proposition.**  $\log n! \sim n \log n$

**Theorem** (Stirling's formula). As  $n \rightarrow \infty$ ,

$$\log \left( \frac{n! e^n}{n^{n+\frac{1}{2}}} \right) = \log \sqrt{2\pi} + O\left(\frac{1}{n}\right)$$

**Corollary.**

$$n! \sim \sqrt{2\pi n} n^{n+\frac{1}{2}} e^{-n}$$

**Proposition** (non-examinable). We use the  $1/12n$  term from the proof above to get a better approximation:

$$\sqrt{2\pi n} n^{n+1/2} e^{-n+\frac{1}{12n+1}} \leq n! \leq \sqrt{2\pi n} n^{n+1/2} e^{-n+\frac{1}{12n}}.$$

## 2 Axioms of probability

### 2.1 Axioms and definitions

**Theorem.**

- (i)  $\mathbb{P}(\emptyset) = 0$
- (ii)  $\mathbb{P}(A^C) = 1 - \mathbb{P}(A)$
- (iii)  $A \subseteq B \Rightarrow \mathbb{P}(A) \leq \mathbb{P}(B)$
- (iv)  $\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cap B)$ .

**Theorem.** If  $A_1, A_2, \dots$  is increasing or decreasing, then

$$\lim_{n \rightarrow \infty} \mathbb{P}(A_n) = \mathbb{P}\left(\lim_{n \rightarrow \infty} A_n\right).$$

### 2.2 Inequalities and formulae

**Theorem** (Boole's inequality). For any  $A_1, A_2, \dots$ ,

$$\mathbb{P}\left(\bigcup_{i=1}^{\infty} A_i\right) \leq \sum_{i=1}^{\infty} \mathbb{P}(A_i).$$

**Theorem** (Inclusion-exclusion formula).

$$\begin{aligned} \mathbb{P}\left(\bigcup_i^n A_i\right) &= \sum_1^n \mathbb{P}(A_i) - \sum_{i_1 < i_2} \mathbb{P}(A_{i_1} \cap A_{i_2}) + \sum_{i_1 < i_2 < i_3} \mathbb{P}(A_{i_1} \cap A_{i_2} \cap A_{i_3}) - \dots \\ &\quad + (-1)^{n-1} \mathbb{P}(A_1 \cap \dots \cap A_n). \end{aligned}$$

**Theorem** (Bonferroni's inequalities). For any events  $A_1, A_2, \dots, A_n$  and  $1 \leq r \leq n$ , if  $r$  is odd, then

$$\begin{aligned} \mathbb{P}\left(\bigcup_1^n A_i\right) &\leq \sum_{i_1} \mathbb{P}(A_{i_1}) - \sum_{i_1 < i_2} \mathbb{P}(A_{i_1} A_{i_2}) + \sum_{i_1 < i_2 < i_3} \mathbb{P}(A_{i_1} A_{i_2} A_{i_3}) + \dots \\ &\quad + \sum_{i_1 < i_2 < \dots < i_r} \mathbb{P}(A_{i_1} A_{i_2} A_{i_3} \dots A_{i_r}). \end{aligned}$$

If  $r$  is even, then

$$\begin{aligned} \mathbb{P}\left(\bigcup_1^n A_i\right) &\geq \sum_{i_1} \mathbb{P}(A_{i_1}) - \sum_{i_1 < i_2} \mathbb{P}(A_{i_1} A_{i_2}) + \sum_{i_1 < i_2 < i_3} \mathbb{P}(A_{i_1} A_{i_2} A_{i_3}) + \dots \\ &\quad - \sum_{i_1 < i_2 < \dots < i_r} \mathbb{P}(A_{i_1} A_{i_2} A_{i_3} \dots A_{i_r}). \end{aligned}$$

### 2.3 Independence

**Proposition.** If  $A$  and  $B$  are independent, then  $A$  and  $B^C$  are independent.

## 2.4 Important discrete distributions

**Theorem** (Poisson approximation to binomial). Suppose  $n \rightarrow \infty$  and  $p \rightarrow 0$  such that  $np = \lambda$ . Then

$$q_k = \binom{n}{k} p^k (1-p)^{n-k} \rightarrow \frac{\lambda^k}{k!} e^{-\lambda}.$$

## 2.5 Conditional probability

**Theorem.**

- (i)  $\mathbb{P}(A \cap B) = \mathbb{P}(A | B)\mathbb{P}(B)$ .
- (ii)  $\mathbb{P}(A \cap B \cap C) = \mathbb{P}(A | B \cap C)\mathbb{P}(B | C)\mathbb{P}(C)$ .
- (iii)  $\mathbb{P}(A | B \cap C) = \frac{\mathbb{P}(A \cap B | C)}{\mathbb{P}(B | C)}$ .
- (iv) The function  $\mathbb{P}(\cdot | B)$  restricted to subsets of  $B$  is a probability function (or measure).

**Proposition.** If  $B_i$  is a partition of the sample space, and  $A$  is any event, then

$$\mathbb{P}(A) = \sum_{i=1}^{\infty} \mathbb{P}(A \cap B_i) = \sum_{i=1}^{\infty} \mathbb{P}(A | B_i)\mathbb{P}(B_i).$$

**Theorem** (Bayes' formula). Suppose  $B_i$  is a partition of the sample space, and  $A$  and  $B_i$  all have non-zero probability. Then for any  $B_i$ ,

$$\mathbb{P}(B_i | A) = \frac{\mathbb{P}(A | B_i)\mathbb{P}(B_i)}{\sum_j \mathbb{P}(A | B_j)\mathbb{P}(B_j)}.$$

Note that the denominator is simply  $\mathbb{P}(A)$  written in a fancy way.

### 3 Discrete random variables

#### 3.1 Discrete random variables

**Theorem.**

- (i) If  $X \geq 0$ , then  $\mathbb{E}[X] \geq 0$ .
- (ii) If  $X \geq 0$  and  $\mathbb{E}[X] = 0$ , then  $\mathbb{P}(X = 0) = 1$ .
- (iii) If  $a$  and  $b$  are constants, then  $\mathbb{E}[a + bX] = a + b\mathbb{E}[X]$ .
- (iv) If  $X$  and  $Y$  are random variables, then  $\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y]$ . This is true even if  $X$  and  $Y$  are not independent.
- (v)  $\mathbb{E}[X]$  is a constant that minimizes  $\mathbb{E}[(X - c)^2]$  over  $c$ .

**Theorem.** For any random variables  $X_1, X_2, \dots, X_n$ , for which the following expectations exist,

$$\mathbb{E} \left[ \sum_{i=1}^n X_i \right] = \sum_{i=1}^n \mathbb{E}[X_i].$$

**Theorem.**

- (i)  $\text{var } X \geq 0$ . If  $\text{var } X = 0$ , then  $\mathbb{P}(X = \mathbb{E}[X]) = 1$ .
- (ii)  $\text{var}(a + bX) = b^2 \text{var}(X)$ . This can be proved by expanding the definition and using the linearity of the expected value.
- (iii)  $\text{var}(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2$ , also proven by expanding the definition.

**Proposition.**

- $\mathbb{E}[I[A]] = \sum_{\omega} p(\omega) I[A](\omega) = \mathbb{P}(A)$ .
- $I[A^C] = 1 - I[A]$ .
- $I[A \cap B] = I[A]I[B]$ .
- $I[A \cup B] = I[A] + I[B] - I[A]I[B]$ .
- $I[A]^2 = I[A]$ .

**Theorem** (Inclusion-exclusion formula).

$$\begin{aligned} \mathbb{P} \left( \bigcup_i^n A_i \right) &= \sum_1^n \mathbb{P}(A_i) - \sum_{i_1 < i_2} \mathbb{P}(A_{i_1} \cap A_{i_2}) + \sum_{i_1 < i_2 < i_3} \mathbb{P}(A_{i_1} \cap A_{i_2} \cap A_{i_3}) - \dots \\ &\quad + (-1)^{n-1} \mathbb{P}(A_1 \cap \dots \cap A_n). \end{aligned}$$

**Theorem.** If  $X_1, \dots, X_n$  are independent random variables, and  $f_1, \dots, f_n$  are functions  $\mathbb{R} \rightarrow \mathbb{R}$ , then  $f_1(X_1), \dots, f_n(X_n)$  are independent random variables.

**Theorem.** If  $X_1, \dots, X_n$  are independent random variables and all the following expectations exists, then

$$\mathbb{E} \left[ \prod X_i \right] = \prod \mathbb{E}[X_i].$$

**Corollary.** Let  $X_1, \dots, X_n$  be independent random variables, and  $f_1, f_2, \dots, f_n$  are functions  $\mathbb{R} \rightarrow \mathbb{R}$ . Then

$$\mathbb{E} \left[ \prod f_i(x_i) \right] = \prod \mathbb{E}[f_i(x_i)].$$

**Theorem.** If  $X_1, X_2, \dots, X_n$  are independent random variables, then

$$\text{var} \left( \sum X_i \right) = \sum \text{var}(X_i).$$

**Corollary.** Let  $X_1, X_2, \dots, X_n$  be independent identically distributed random variables (iid rvs). Then

$$\text{var} \left( \frac{1}{n} \sum X_i \right) = \frac{1}{n} \text{var}(X_1).$$

### 3.2 Inequalities

**Proposition.** If  $f$  is differentiable and  $f''(x) \geq 0$  for all  $x \in (a, b)$ , then it is convex. It is strictly convex if  $f''(x) > 0$ .

**Theorem** (Jensen's inequality). If  $f : (a, b) \rightarrow \mathbb{R}$  is convex, then

$$\sum_{i=1}^n p_i f(x_i) \geq f \left( \sum_{i=1}^n p_i x_i \right)$$

for all  $p_1, p_2, \dots, p_n$  such that  $p_i \geq 0$  and  $\sum p_i = 1$ , and  $x_i \in (a, b)$ .

This says that  $\mathbb{E}[f(X)] \geq f(\mathbb{E}[X])$  (where  $\mathbb{P}(X = x_i) = p_i$ ).

If  $f$  is strictly convex, then equalities hold only if all  $x_i$  are equal, i.e.  $X$  takes only one possible value.

**Corollary** (AM-GM inequality). Given  $x_1, \dots, x_n$  positive reals, then

$$\left( \prod x_i \right)^{1/n} \leq \frac{1}{n} \sum x_i.$$

**Theorem** (Cauchy-Schwarz inequality). For any two random variables  $X, Y$ ,

$$(\mathbb{E}[XY])^2 \leq \mathbb{E}[X^2]\mathbb{E}[Y^2].$$

**Theorem** (Markov inequality). If  $X$  is a random variable with  $\mathbb{E}|X| < \infty$  and  $\varepsilon > 0$ , then

$$\mathbb{P}(|X| \geq \varepsilon) \leq \frac{\mathbb{E}|X|}{\varepsilon}.$$

**Theorem** (Chebyshev inequality). If  $X$  is a random variable with  $\mathbb{E}[X^2] < \infty$  and  $\varepsilon > 0$ , then

$$\mathbb{P}(|X| \geq \varepsilon) \leq \frac{\mathbb{E}[X^2]}{\varepsilon^2}.$$



### 3.3 Weak law of large numbers

**Theorem** (Weak law of large numbers). Let  $X_1, X_2, \dots$  be iid random variables, with mean  $\mu$  and var  $\sigma^2$ .

Let  $S_n = \sum_{i=1}^n X_i$ .

Then for all  $\varepsilon > 0$ ,

$$\mathbb{P}\left(\left|\frac{S_n}{n} - \mu\right| \geq \varepsilon\right) \rightarrow 0$$

as  $n \rightarrow \infty$ .

We say,  $\frac{S_n}{n}$  tends to  $\mu$  (in probability), or

$$\frac{S_n}{n} \rightarrow_p \mu.$$

**Theorem** (Strong law of large numbers).

$$\mathbb{P}\left(\frac{S_n}{n} \rightarrow \mu \text{ as } n \rightarrow \infty\right) = 1.$$

We say

$$\frac{S_n}{n} \rightarrow_{\text{as}} \mu,$$

where “as” means “almost surely”.

### 3.4 Multiple random variables

**Proposition.**

- (i)  $\text{cov}(X, c) = 0$  for constant  $c$ .
- (ii)  $\text{cov}(X + c, Y) = \text{cov}(X, Y)$ .
- (iii)  $\text{cov}(X, Y) = \text{cov}(Y, X)$ .
- (iv)  $\text{cov}(X, Y) = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$ .
- (v)  $\text{cov}(X, X) = \text{var}(X)$ .
- (vi)  $\text{var}(X + Y) = \text{var}(X) + \text{var}(Y) + 2\text{cov}(X, Y)$ .
- (vii) If  $X, Y$  are independent,  $\text{cov}(X, Y) = 0$ .

**Proposition.**  $|\text{corr}(X, Y)| \leq 1$ .

**Theorem.** If  $X$  and  $Y$  are independent, then

$$\mathbb{E}[X | Y] = \mathbb{E}[X]$$

**Theorem** (Tower property of conditional expectation).

$$\mathbb{E}_Y[\mathbb{E}_X[X | Y]] = \mathbb{E}_X[X],$$

where the subscripts indicate what variable the expectation is taken over.

### 3.5 Probability generating functions

**Theorem.** The distribution of  $X$  is uniquely determined by its probability generating function.

**Theorem** (Abel's lemma).

$$\mathbb{E}[X] = \lim_{z \rightarrow 1} p'(z).$$

If  $p'(z)$  is continuous, then simply  $\mathbb{E}[X] = p'(1)$ .

**Theorem.**

$$\mathbb{E}[X(X-1)] = \lim_{z \rightarrow 1} p''(z).$$

**Theorem.** Suppose  $X_1, X_2, \dots, X_n$  are independent random variables with pgfs  $p_1, p_2, \dots, p_n$ . Then the pgf of  $X_1 + X_2 + \dots + X_n$  is  $p_1(z)p_2(z) \cdots p_n(z)$ .

## 4 Interesting problems

### 4.1 Branching processes

**Theorem.**

$$F_{n+1}(z) = F_n(F(z)) = F(F(F(\dots F(z)\dots))) = F(F_n(z)).$$

**Theorem.** Suppose

$$\mathbb{E}[X_1] = \sum k p_k = \mu$$

and

$$\text{var}(X_1) = \mathbb{E}[(X - \mu)^2] = \sum (k - \mu)^2 p_k < \infty.$$

Then

$$\mathbb{E}[X_n] = \mu^n, \quad \text{var } X_n = \sigma^2 \mu^{n-1} (1 + \mu + \mu^2 + \dots + \mu^{n-1}).$$

**Theorem.** The probability of extinction  $q$  is the smallest root to the equation  $q = F(q)$ . Write  $\mu = \mathbb{E}[X_1]$ . Then if  $\mu \leq 1$ , then  $q = 1$ ; if  $\mu > 1$ , then  $q < 1$ .

### 4.2 Random walk and gambler's ruin

## 5 Continuous random variables

### 5.1 Continuous random variables

**Proposition.** The exponential random variable is *memoryless*, i.e.

$$\mathbb{P}(X \geq x + z \mid X \geq x) = \mathbb{P}(X \geq z).$$

This means that, say if  $X$  measures the lifetime of a light bulb, knowing it has already lasted for 3 hours does not give any information about how much longer it will last.

**Theorem.** If  $X$  is a continuous random variable, then

$$\mathbb{E}[X] = \int_0^{\infty} \mathbb{P}(X \geq x) \, dx - \int_0^{\infty} \mathbb{P}(X \leq -x) \, dx.$$

### 5.2 Stochastic ordering and inspection paradox

### 5.3 Jointly distributed random variables

**Theorem.** If  $X$  and  $Y$  are jointly continuous random variables, then they are individually continuous random variables.

**Proposition.** For independent continuous random variables  $X_i$ ,

- (i)  $\mathbb{E}[\prod X_i] = \prod \mathbb{E}[X_i]$
- (ii)  $\text{var}(\sum X_i) = \sum \text{var}(X_i)$

### 5.4 Geometric probability

### 5.5 The normal distribution

**Proposition.**

$$\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(x-\mu)^2} \, dx = 1.$$

**Proposition.**  $\mathbb{E}[X] = \mu$ .

**Proposition.**  $\text{var}(X) = \sigma^2$ .

### 5.6 Transformation of random variables

**Theorem.** If  $X$  is a continuous random variable with a pdf  $f(x)$ , and  $h(x)$  is a continuous, strictly increasing function with  $h^{-1}(x)$  differentiable, then  $Y = h(X)$  is a random variable with pdf

$$f_Y(y) = f_X(h^{-1}(y)) \frac{d}{dy} h^{-1}(y).$$

**Theorem.** Let  $U \sim U[0, 1]$ . For any strictly increasing distribution function  $F$ , the random variable  $X = F^{-1}U$  has distribution function  $F$ .

**Proposition.**  $(Y_1, \dots, Y_n)$  has density

$$g(y_1, \dots, y_n) = f(s_1(y_1, \dots, y_n), \dots, s_n(y_1, \dots, y_n)) |J|$$

if  $(y_1, \dots, y_n) \in S$ , 0 otherwise.

## 5.7 Moment generating functions

**Theorem.** The mgf determines the distribution of  $X$  provided  $m(\theta)$  is finite for all  $\theta$  in some interval containing the origin.

**Theorem.** The  $r$ th moment  $X$  is the coefficient of  $\frac{\theta^r}{r!}$  in the power series expansion of  $m(\theta)$ , and is

$$\mathbb{E}[X^r] = \left. \frac{d^r}{d\theta^r} m(\theta) \right|_{\theta=0} = m^{(r)}(0).$$

**Theorem.** If  $X$  and  $Y$  are independent random variables with moment generating functions  $m_X(\theta), m_Y(\theta)$ , then  $X + Y$  has mgf  $m_{X+Y}(\theta) = m_X(\theta)m_Y(\theta)$ .

## 6 More distributions

### 6.1 Cauchy distribution

**Proposition.** The mean of the Cauchy distribution is undefined, while  $\mathbb{E}[X^2] = \infty$ .

### 6.2 Gamma distribution

### 6.3 Beta distribution\*

### 6.4 More on the normal distribution

**Proposition.** The moment generating function of  $N(\mu, \sigma^2)$  is

$$\mathbb{E}[e^{\theta X}] = \exp\left(\theta\mu + \frac{1}{2}\theta^2\sigma^2\right).$$

**Theorem.** Suppose  $X, Y$  are independent random variables with  $X \sim N(\mu_1, \sigma_1^2)$ , and  $Y \sim N(\mu_2, \sigma_2^2)$ . Then

- (i)  $X + Y \sim N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$ .
- (ii)  $aX \sim N(a\mu_1, a^2\sigma_1^2)$ .

### 6.5 Multivariate normal

## 7 Central limit theorem

**Theorem** (Central limit theorem). Let  $X_1, X_2, \dots$  be iid random variables with  $\mathbb{E}[X_i] = \mu$ ,  $\text{var}(X_i) = \sigma^2 < \infty$ . Define

$$S_n = X_1 + \dots + X_n.$$

Then for all finite intervals  $(a, b)$ ,

$$\lim_{n \rightarrow \infty} \mathbb{P} \left( a \leq \frac{S_n - n\mu}{\sigma\sqrt{n}} \leq b \right) = \int_a^b \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}t^2} dt.$$

Note that the final term is the pdf of a standard normal. We say

$$\frac{S_n - n\mu}{\sigma\sqrt{n}} \rightarrow_D N(0, 1).$$

**Theorem** (Continuity theorem). If the random variables  $X_1, X_2, \dots$  have mgf's  $m_1(\theta), m_2(\theta), \dots$  and  $m_n(\theta) \rightarrow m(\theta)$  as  $n \rightarrow \infty$  for all  $\theta$ , then  $X_n \rightarrow_D$  the random variable with mgf  $m(\theta)$ .

## **8 Summary of distributions**

### **8.1 Discrete distributions**

### **8.2 Continuous distributions**